

Mod- ϕ convergence II: dependency graphs

Valentin Féray

(joint work with Pierre-Loïc Méliot and Ashkan Nikeghbali)

Institut für Mathematik, Universität Zürich

Summer school in Villa Volpi, Lago Maggiore,
Aug. 31st - Sep 7th, 2017



Universität
Zürich^{UZH}

Reminder

Definition (uniform control on cumulants)

A sequence (S_n) admits a **uniform control on cumulants** with DNA (D_n, N_n, A) and limits σ^2 and L if $D_n = o(N_n)$, $N_n \rightarrow +\infty$ and

$$\forall r \geq 2, |\kappa^{(r)}(S_n)| \leq N_n (2D_n)^{r-1} r^{r-2} A^r;$$
$$\frac{\kappa^{(2)}(S_n)}{N_n D_n} = (\sigma_n)^2 \xrightarrow{n \rightarrow \infty} \sigma^2; \quad \frac{\kappa^{(3)}(S_n)}{N_n (D_n)^2} = L_n \xrightarrow{n \rightarrow \infty} L.$$

This yields **mod-Gaussian convergence** of a suited renormalization of S_n , hence **deviation probability estimate** and a **bound on the speed on convergence** in a CLT.

Today's talk

See a setting where the inequality above holds: [dependency graphs](#) and a weighted variant.

[Many applications](#): subword/subgraph counts in random words/graphs, magnetization in Ising model, linear statistics of Markov chains.

We will also discuss a weaker framework, which is useful to prove CLT, but where we cannot prove mod-Gaussian convergence.

Dependency graphs

Definition (Malyshev, '80, Petrovskaya/Leontovich, '82, Janson, '88)

A graph L with vertex set A is a dependency graph for the family $\{Y_\alpha, \alpha \in A\}$ if

- if A_1 and A_2 are disconnected subsets in L , then $\{Y_\alpha, \alpha \in A_1\}$ and $\{Y_\alpha, \alpha \in A_2\}$ are independent.

Roughly: there are edges between pairs of **dependent** random variables.

Dependency graphs

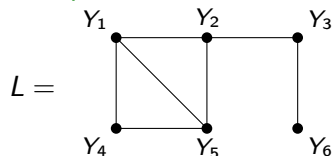
Definition (Malyshev, '80, Petrovskaya/Leontovich, '82, Janson, '88)

A graph L with vertex set A is a dependency graph for the family $\{Y_\alpha, \alpha \in A\}$ if

- if A_1 and A_2 are disconnected subsets in L , then $\{Y_\alpha, \alpha \in A_1\}$ and $\{Y_\alpha, \alpha \in A_2\}$ are independent.

Roughly: there are edges between pairs of **dependent** random variables.

Example 1



Y_2 and Y_4 are independent;
 $\{Y_1, Y_4, Y_5\}$ independent from $\{Y_3, Y_6\}$.

Dependency graphs

Definition (Malyshev, '80, Petrovskaya/Leontovich, '82, Janson, '88)

A graph L with vertex set A is a dependency graph for the family $\{Y_\alpha, \alpha \in A\}$ if

- if A_1 and A_2 are disconnected subsets in L , then $\{Y_\alpha, \alpha \in A_1\}$ and $\{Y_\alpha, \alpha \in A_2\}$ are independent.

Roughly: there are edges between pairs of **dependent** random variables.

Example 2 (triangles in $G(n, p)$)

Consider $G = G(n, p)$. Let $A = \{\Delta \in \binom{[n]}{3}\}$ (set of potential triangles) and

$\{\Delta_1, \Delta_2\} \in E_L$ iff Δ_1 and Δ_2 share an edge in G .

Then L is a dependency graph for the family $\{\mathbf{1}_{\Delta \subset G}, \Delta \in \binom{[n]}{3}\}$.

Dependency graphs

Definition (Malyshev, '80, Petrovskaya/Leontovich, '82, Janson, '88)

A graph L with vertex set A is a dependency graph for the family $\{Y_\alpha, \alpha \in A\}$ if

- if A_1 and A_2 are disconnected subsets in L , then $\{Y_\alpha, \alpha \in A_1\}$ and $\{Y_\alpha, \alpha \in A_2\}$ are independent.

Roughly: there are edges between pairs of **dependent** random variables.

Example 2 (triangles in $G(n, p)$)

Note: L has degree $\mathcal{O}(n)$

Consider $G = G(n, p)$. Let $A = \{\Delta \in \binom{[n]}{3}\}$ (set of potential triangles) and

$\{\Delta_1, \Delta_2\} \in E_L$ iff Δ_1 and Δ_2 share an edge in G .

Then L is a dependency graph for the family $\{\mathbf{1}_{\Delta \subset G}, \Delta \in \binom{[n]}{3}\}$.

Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < A$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $S_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(S_n)$.

Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < A$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $S_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(S_n)$.

Theorem (Janson, 1988)

Assume that $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some integer s . Then S_n satisfies a CLT.

Janson's normality criterion

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < A$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $S_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(S_n)$.

Theorem (Janson, 1988)

Assume that $\left(\frac{N_n}{D_n}\right)^{1/s} \frac{D_n}{\sigma_n} \rightarrow 0$ for some integer s . Then S_n satisfies a CLT.

For triangles, $N_n = \binom{n}{3}$, $D_n = O(n)$, while $\sigma_n \asymp n^2$. (for fixed p)

Corollary

Fix p in $(0, 1)$. Then T_n satisfies a CLT.

(also true for $p_n \rightarrow 0$ with $np_n \rightarrow \infty$; originally proved by Rucinski, 1988).

Applications of dependency graphs to CLT results

- mathematical modelization of cell populations (Petrovskaya, Leontovich, 82);
- subgraph counts in random graphs (Janson, Baldi, Rinott, Penrose, 88, 89, 95, 03);
- Geometric probability (Avram, Bertsimas, Penrose, Yukich, Bárány, Vu, 93, 05, 07);
- pattern occurrences in random permutations (Bóna, Janson, Hitchenko, Nakamura, Zeilberger, 07, 09, 14).
- m -dependence (Hoeffding, Robbins, 53, ...; now widely used in statistics) is a special case.

(Some of these applications use variants of Janson's normality criterion, which are more technical to state and omitted here. . .)

A bound for cumulants

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < A$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $S_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(S_n)$.

A bound for cumulants

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < A$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $S_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(S_n)$.

Lemma (Janson, 1988)

$$\kappa_r(S_n) \leq C_r N_n D_n^{r-1} A^r,$$

for some universal constant C_r .

A bound for cumulants

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < A$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $S_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(S_n)$.

Lemma (Janson, 1988)

$$\kappa_r(S_n) \leq C_r N_n D_n^{r-1} A^r,$$

for some universal constant C_r .

Döring and Eichelsbacher, 2012: one can take $C_r = (2e)^r (r!)^3$.

A bound for cumulants

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < A$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $S_n = \sum_{i=1}^{N_n} Y_{n,i}$ and $\sigma_n^2 = \text{Var}(S_n)$.

Lemma (Janson, 1988)

$$\kappa_r(S_n) \leq C_r N_n D_n^{r-1} A^r,$$

for some universal constant C_r .

Döring and Eichelsbacher, 2012: one can take $C_r = (2e)^r (r!)^3$.

FMN, 2016: one can take $C_r = 2^{r-1} r^{r-2}$,

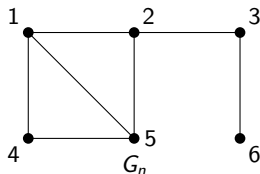
i.e. we have the uniform bound on cumulants.

A bound for mixed cumulants

How to bound $\kappa_r(S_n)$?

$$\kappa_r(S_n) = \kappa(S_n, \dots, S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n,i_1}, \dots, Y_{n,i_r})$$

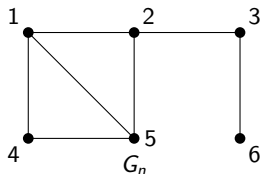
$\kappa(Y_{n,i_1}, \dots, Y_{n,i_r})$ is the mixed cumulants; multilinear version of cumulants.



A bound for mixed cumulants

How to bound $\kappa_r(S_n)$?

$$\kappa_r(S_n) = \kappa(S_n, \dots, S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n,i_1}, \dots, Y_{n,i_r})$$



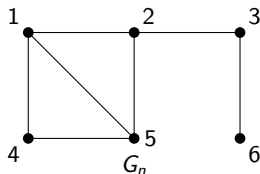
Most terms are zero: $\kappa(Y_{n,i_1}, \dots, Y_{n,i_r}) = 0$ unless the induced graph $G_n[\{i_1, \dots, i_r\}]$ is connected.

e.g. $\kappa(Y_{n,1}, Y_{n,3}, Y_{n,4}, Y_{n,5}) = 0$

A bound for mixed cumulants

How to bound $\kappa_r(S_n)$?

$$\kappa_r(S_n) = \kappa(S_n, \dots, S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n, i_1}, \dots, Y_{n, i_r})$$



Most terms are zero: $\kappa(Y_{n, i_1}, \dots, Y_{n, i_r}) = 0$ unless the induced graph $G_n[\{i_1, \dots, i_r\}]$ is connected.

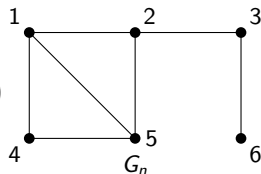
Usual strategy: bound each term $\kappa(Y_{n, i_1}, \dots, Y_{n, i_r})$ and the number of non-zero terms.

We prove a bound depending on the induced graph $G_n[\{i_1, \dots, i_r\}]$.

A bound for mixed cumulants

How to bound $\kappa_r(S_n)$?

$$\kappa_r(S_n) = \kappa(S_n, \dots, S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n, i_1}, \dots, Y_{n, i_r})$$



Most terms are zero: $\kappa(Y_{n, i_1}, \dots, Y_{n, i_r}) = 0$ unless the induced graph $G_n[\{i_1, \dots, i_r\}]$ is connected.

Proposition (FMN, 2016)

$$|\kappa(Y_{n, i_1}, \dots, Y_{n, i_r})| \leq 2^{\ell-1} A^r \text{ST}(G_n[\{i_1, \dots, i_r\}]),$$

where $\text{ST}(G)$ denotes the number of spanning tree of a graph G .

e.g. $|\kappa(Y_{n,1}, Y_{n,2}, Y_{n,4}, Y_{n,5})| \leq 2^3 A^4 8;$

$$|\kappa(Y_{n,1}, Y_{n,2}, Y_{n,3}, Y_{n,4})| \leq 2^3 A^4 1.$$

From mixed cumulants to cumulants

Recall that $\kappa_r(S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n, i_1}, \dots, Y_{n, i_r})$.

Thus

$$|\kappa_r(S_n)| \leq \sum_{i_1, \dots, i_r} 2^{r-1} A^r \text{ST}(G_n[\{i_1, \dots, i_r\}]).$$

From mixed cumulants to cumulants

Recall that $\kappa_r(S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n, i_1}, \dots, Y_{n, i_r})$.

Thus

$$|\kappa_r(S_n)| \leq 2^{r-1} A^r \sum_{T \text{ Cayley tree}} |\{(i_1, \dots, i_r \text{ s.t. } T \subset G_n[\{i_1, \dots, i_r\}])\}|.$$

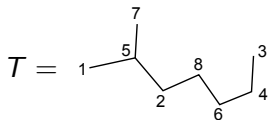
From mixed cumulants to cumulants

Recall that $\kappa_r(S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n, i_1}, \dots, Y_{n, i_r})$.

Thus

$$|\kappa_r(S_n)| \leq 2^{r-1} A^r \sum_{T \text{ Cayley tree}} |\{(i_1, \dots, i_r \text{ s.t. } T \subset G_n[\{i_1, \dots, i_r\}]\}|.$$

Fix a Cayley tree T . For how many lists i_1, \dots, i_r is it contained in $G_n[\{i_1, \dots, i_r\}]$?



From mixed cumulants to cumulants

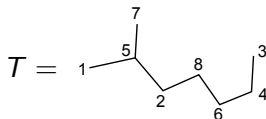
Recall that $\kappa_r(S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n, i_1}, \dots, Y_{n, i_r})$.

Thus

$$|\kappa_r(S_n)| \leq 2^{r-1} A^r \sum_{T \text{ Cayley tree}} |\{(i_1, \dots, i_r \text{ s.t. } T \subset G_n[\{i_1, \dots, i_r\}]\}|.$$

Fix a Cayley tree T . For how many lists i_1, \dots, i_r is it contained in $G_n[\{i_1, \dots, i_r\}]$?

- Choose any i_1 : N_n choices ;



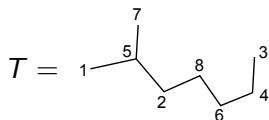
From mixed cumulants to cumulants

Recall that $\kappa_r(S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n,i_1}, \dots, Y_{n,i_r})$.

Thus

$$|\kappa_r(S_n)| \leq 2^{r-1} A^r \sum_{T \text{ Cayley tree}} |\{(i_1, \dots, i_r \text{ s.t. } T \subset G_n[\{i_1, \dots, i_r\}]\}|.$$

Fix a Cayley tree T . For how many lists i_1, \dots, i_r is it contained in $G_n[\{i_1, \dots, i_r\}]$?



- Choose any i_1 : N_n choices ;
- i_5 should be a neighbour of i_1 in G_n (or i_1 itself): D_n choices ;

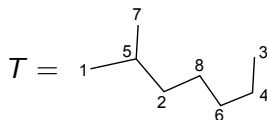
From mixed cumulants to cumulants

Recall that $\kappa_r(S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n, i_1}, \dots, Y_{n, i_r})$.

Thus

$$|\kappa_r(S_n)| \leq 2^{r-1} A^r \sum_{T \text{ Cayley tree}} |\{(i_1, \dots, i_r \text{ s.t. } T \subset G_n[\{i_1, \dots, i_r\}]\}|.$$

Fix a Cayley tree T . For how many lists i_1, \dots, i_r is it contained in $G_n[\{i_1, \dots, i_r\}]$?



- Choose any i_1 : N_n choices ;
- i_5 should be a neighbour of i_1 in G_n (or i_1 itself): D_n choices ;
- i_2 should be a neighbour of i_5 in G_n (or i_5 itself). Also D_n choices.
- ...

From mixed cumulants to cumulants

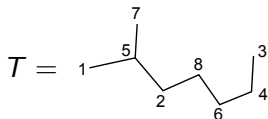
Recall that $\kappa_r(S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n, i_1}, \dots, Y_{n, i_r})$.

Thus

$$|\kappa_r(S_n)| \leq 2^{r-1} A^r \sum_{T \text{ Cayley tree}} |\{(i_1, \dots, i_r \text{ s.t. } T \subset G_n[\{i_1, \dots, i_r\}]\}|.$$

Fix a Cayley tree T . For how many lists i_1, \dots, i_r is it contained in $G_n[\{i_1, \dots, i_r\}]$?

$$N_n D_n^{r-1}$$



- Choose any i_1 : N_n choices ;
- i_5 should be a neighbour of i_1 in G_n (or i_1 itself): D_n choices ;
- i_2 should be a neighbour of i_5 in G_n (or i_5 itself). Also D_n choices.
- ...

From mixed cumulants to cumulants

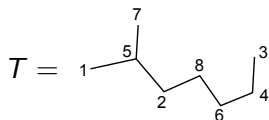
Recall that $\kappa_r(S_n) = \sum_{i_1, \dots, i_r} \kappa(Y_{n, i_1}, \dots, Y_{n, i_r})$.

Thus

$$|\kappa_r(S_n)| \leq 2^{r-1} A^r \sum_{T \text{ Cayley tree}} |\{(i_1, \dots, i_r \text{ s.t. } T \subset G_n[\{i_1, \dots, i_r\}]\}|.$$

Fix a Cayley tree T . For how many lists i_1, \dots, i_r is it contained in $G_n[\{i_1, \dots, i_r\}]$?

$$N_n D_n^{r-1}$$



- Choose any i_1 : N_n choices ;
- i_5 should be a neighbour of i_1 in G_n (or i_1 itself): D_n choices ;
- i_2 should be a neighbour of i_5 in G_n (or i_5 itself). Also D_n choices.
- ... $|\kappa_r(S_n)| \leq 2^{r-1} r^{r-2} N_n D_n^{r-1} A^r$

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < A$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $S_n = \sum_{i=1}^{N_n} Y_{n,i}$, $\sigma_n^2 = \text{Var}(S_n)$ and $L_n^3 = \kappa^3(S_n)$.

Setting: for each n ,

- $\{Y_{n,i}, 1 \leq i \leq N_n\}$ is a family of bounded random variables; $|Y_{n,i}| < A$ a.s.
- we have a dependency graph L_n with maximal degree $D_n - 1$.
- we set $S_n = \sum_{i=1}^{N_n} Y_{n,i}$, $\sigma_n^2 = \text{Var}(S_n)$ and $L_n^3 = \kappa^3(S_n)$.

Theorem (FMN, 2017)

Assume that $\frac{\sigma_n^2}{N_n D_n} \rightarrow \sigma^2 > 0$. Then the error term in the CLT for $\frac{1}{\sigma_n}(S_n - \mathbb{E}(S_n))$ is $O(\sqrt{D_n/N_n})$.

If furthermore $\frac{\sigma_n^2}{N_n D_n} \rightarrow L^3 \neq 0$, the normality zone is $o((D_n/N_n)^{1/6})$.

Proof: uniform bound of cumulants + general results from yesterday.

The speed of convergence was already known from Rinott (1994), through Stein's method.

Application 1: subwords

Consider a random word with i.i.d. letters

$$w = w_1 \dots w_N$$

We are interested in the number of (non-necessarily consecutive) occurrences of a given word u (e.g., for $u = aab$, we count triplet of indices (i, j, k) with $i < j < k$ and $w_i = w_j = a, w_k = b$).

A CLT proved by Flajolet, Szpankowski and Vallée, '01.

Proposition

The number of copies of a given word of length k in a random word with i.i.d. letters admits a uniform control on cumulants with DNA $(n^{k-1}, n^k, 1)$.

Application 1: subwords

Consider a random word with i.i.d. letters

$$w = w_1 \dots w_N$$

We are interested in the number of (non-necessarily consecutive) occurrences of a given word u (e.g., for $u = aab$, we count triplet of indices (i, j, k) with $i < j < k$ and $w_i = w_j = a, w_k = b$).

A CLT proved by Flajolet, Szpankowski and Vallée, '01.

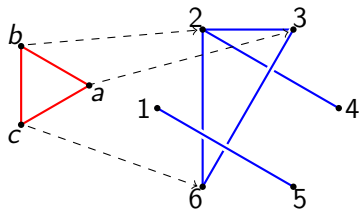
Proposition

The number of copies of a given word of length k in a random word with i.i.d. letters admits a uniform control on cumulants with DNA $(n^{k-1}, n^k, 1)$.

Proof: bounds on cumulant given by an obvious dependency graph, estimate for the second and third cumulants given by a combinatorial expansion.

Application 2: subgraph counts

Copies of F in a random graph G

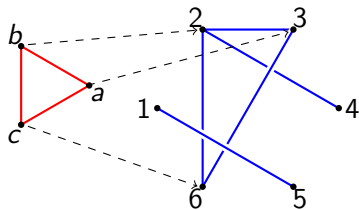


Proposition

The number of copies of a fixed F in $G(n, p)$ (p fixed) admits a uniform control on cumulants with DNA $(n^{|V_G| - 2}, n^{|V_G|}, 1)$ and $\sigma^2 > 0$.

Application 2: subgraph counts

Copies of F in a random graph G



Proposition

The number of copies of a fixed F in $G(n, p)$ (p fixed) admits a uniform control on cumulants with DNA $(n^{|V_G|-2}, n^{|V_G|}, 1)$ and $\sigma^2 > 0$.

Proof: as in subword case.

☹ When $p = p_n \rightarrow 0$, we do not have a good uniform control on cumulants.

We will now discuss a weighted analogue of dependency graphs.

Goal: consider sum of **pairwise dependent** random variables, where the dependencies are **asymptotically small**.

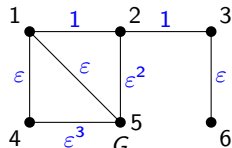
Uniform weighted dependency graph

Let $S = \sum_{i=1}^N Y_i$ as above and let G be an edge weighted graph with vertex set $[N]$ (weights are in $[0, 1]$).

Proposition

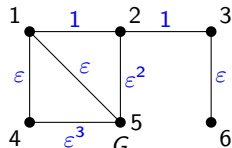
Assume that, for any i_1, \dots, i_r in $[N]$, we have

$$|\kappa(Y_{i_1}, \dots, Y_{i_r})| \leq C^r \sum_{T \text{ Cayley tree}} \prod_{\{j,k\} \in E_T} w(\{i_j, i_k\}), \quad (\text{UWDG})$$



Uniform weighted dependency graph

Let $S = \sum_{i=1}^N Y_i$ as above and let G be an edge weighted graph with vertex set $[N]$ (weights are in $[0, 1]$).



Proposition

Assume that, for any i_1, \dots, i_r in $[N]$, we have

$$|\kappa(Y_{i_1}, \dots, Y_{i_r})| \leq C^r \sum_{T \text{ Cayley tree}} \prod_{\{j,k\} \in E_T} w(\{i_j, i_k\}), \quad (\text{UWDG})$$

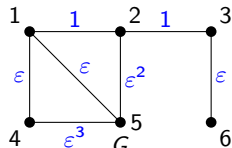
e.g. $\kappa(Y_{n,1}, Y_{n,3}, Y_{n,4}, Y_{n,5}) = 0;$

$$|\kappa(Y_{n,1}, Y_{n,2}, Y_{n,3}, Y_{n,4})| \leq C^4 \varepsilon;$$

$$|\kappa(Y_{n,1}, Y_{n,2}, Y_{n,4}, Y_{n,5})| \leq C^4 (\varepsilon^2 + \varepsilon^3 + 3\varepsilon^4 + \varepsilon^5 + 2\varepsilon^6);$$

Uniform weighted dependency graph

Let $S = \sum_{i=1}^N Y_i$ as above and let G be an edge weighted graph with vertex set $[N]$ (weights are in $[0, 1]$).



Proposition

Assume that, for any i_1, \dots, i_r in $[N]$, we have

$$|\kappa(Y_{i_1}, \dots, Y_{i_r})| \leq C^r \sum_{T \text{ Cayley tree}} \prod_{\{j,k\} \in E_T} w(\{i_j, i_k\}), \quad (\text{UWDG})$$

then $|\kappa_r(S)| \leq C^r N D^{r-1}$, where $D - 1$ is the maximal *weighted degree* of the graph.

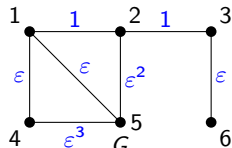
$$\text{e.g. } \kappa(Y_{n,1}, Y_{n,3}, Y_{n,4}, Y_{n,5}) = 0;$$

$$|\kappa(Y_{n,1}, Y_{n,2}, Y_{n,3}, Y_{n,4})| \leq C^4 \varepsilon;$$

$$|\kappa(Y_{n,1}, Y_{n,2}, Y_{n,4}, Y_{n,5})| \leq C^4 (\varepsilon^2 + \varepsilon^3 + 3\varepsilon^4 + \varepsilon^5 + 2\varepsilon^6);$$

Uniform weighted dependency graph

Let $S = \sum_{i=1}^N Y_i$ as above and let G be an edge weighted graph with vertex set $[N]$ (weights are in $[0, 1]$).



Proposition

Assume that, for any i_1, \dots, i_r in $[N]$, we have

$$|\kappa(Y_{i_1}, \dots, Y_{i_r})| \leq C^r \sum_{T \text{ Cayley tree}} \prod_{\{j,k\} \in E_T} w(\{i_j, i_k\}), \quad (\text{UWDG})$$

then $|\kappa_r(S)| \leq C^r N D^{r-1}$, where $D - 1$ is the maximal *weighted degree* of the graph.

Proof: simple adaptation of the case of dependency graphs (which corresponds to edges of G having weights 0 or 1).

When (UWDG) holds, we say that G is a C -uniform weighted dependency graph for $\{Y_i, 1 \leq i \leq N\}$.

Linear statistics in Markov chains

Setting:

- Let $(M_i)_{i \geq 0}$ be an irreducible aperiodic Markov chain on a finite space state S ;
- Assume M_0 is distributed with the stationary distribution π ;
- Set $Y_i = f_i(M_i)$ for some functions f_i uniformly bounded by B .

Linear statistics in Markov chains

Setting:

- Let $(M_i)_{i \geq 0}$ be an irreducible aperiodic Markov chain on a finite space state S ;
- Assume M_0 is distributed with the stationary distribution π ;
- Set $Y_i = f_i(M_i)$ for some functions f_i uniformly bounded by B .

Proposition (FMN, 2017, based on Saulis, Stateljivčius, 1991)

There exists $\varepsilon > 0$ depending on the transition matrix such that the complete graph on $\mathbb{N}_{\geq 0}$ with weights $w(\{s, t\}) = \varepsilon^{|s-t|}$ is a $4B$ -uniform weighted dependency graph for the Y_i 's.

The maximal weighted degree of the restriction to $[n]$ is constant!

Linear statistics in Markov chains

Setting:

- Let $(M_i)_{i \geq 0}$ be an irreducible aperiodic Markov chain on a finite space state S ;
- Assume M_0 is distributed with the stationary distribution π ;
- Set $Y_i = f_i(M_i)$ for some functions f_i uniformly bounded by B .

Proposition (FMN, 2017, based on Saulis, Stalėivičius, 1991)

There exists $\varepsilon > 0$ depending on the transition matrix such that the complete graph on $\mathbb{N}_{\geq 0}$ with weights $w(\{s, t\}) = \varepsilon^{|s-t|}$ is a $4B$ -uniform weighted dependency graph for the Y_i 's.

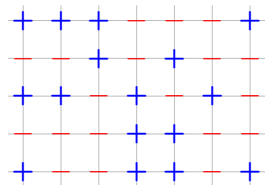
The maximal weighted degree of the restriction to $[n]$ is constant!

→ deviation and speed of convergence estimates for linear statistics of Markov chains (for speed of convergence, see Bolthausen, 1980).

Ising model

Statistical mechanics model to modelize ferromagnetism.

$$P(\omega) \propto \exp \left(\beta \sum_{i \sim j} \omega_i \omega_j + h \sum_i \omega_i \right)$$

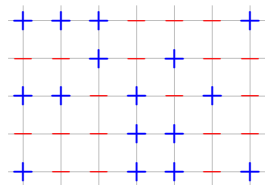


Defined a priori for finite subsets $\Lambda \subseteq \mathbb{Z}^d$, but we can take the “thermodynamic limit” $\Lambda \uparrow \mathbb{Z}^d$ (with + boundary conditions).

Ising model

Statistical mechanics model to modelize ferromagnetism.

$$P(\omega) \propto \exp \left(\beta \sum_{i \sim j} \omega_i \omega_j + h \sum_i \omega_i \right)$$



Proposition (Duneau, Jagolnitzer, Souillard, 1974)

In the thermodynamic limit, for $h \neq 0$ or $h = 0$ and sufficiently small β , there exists $\varepsilon(d), C(d) > 0$ such that the complete graph on $\mathbb{N}_{\geq 0}$ with weights $w(\{s, t\}) = \varepsilon(d)^{\|s-t\|}$ is a $C(d)$ -uniform weighted dependency graph for the σ_i 's.

In particular, for disjoint x_1, \dots, x_r ,

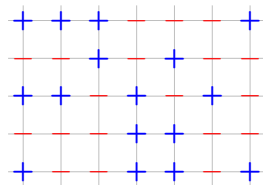
$$\kappa(\sigma_{x_1}, \dots, \sigma_{x_r}) = \mathcal{O}_r(\varepsilon^{\ell_T(x_1, \dots, x_r)}),$$

where $\ell_T(x_1, \dots, x_r)$ is the smallest length of a tree connecting x_1, \dots, x_r .

Ising model

Statistical mechanics model to modelize ferromagnetism.

$$P(\omega) \propto \exp \left(\beta \sum_{i \sim j} \omega_i \omega_j + h \sum_i \omega_i \right)$$



Proposition (Duneau, Jagolnitzer, Souillard, 1974)

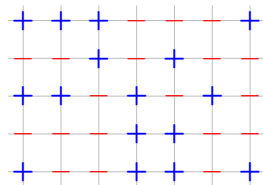
*In the thermodynamic limit, for $h \neq 0$ or $h = 0$ and sufficiently small β , there exists $\varepsilon(d), C(d) > 0$ such that **the complete graph on $\mathbb{N}_{\geq 0}$ with weights $w(\{s, t\}) = \varepsilon(d)^{\|s-t\|}$ is a $C(d)$ -uniform weighted dependency graph for the σ_i 's.***

The same **cannot** be true for large β ($\beta > \beta_c(d)$).

Ising model

Statistical mechanics model to modelize ferromagnetism.

$$P(\omega) \propto \exp \left(\beta \sum_{i \sim j} \omega_i \omega_j + h \sum_i \omega_i \right)$$



Proposition (Duneau, Jagolnitzer, Souillard, 1974)

In the thermodynamic limit, for $h \neq 0$ or $h = 0$ and sufficiently small β , there exists $\varepsilon(d), C(d) > 0$ such that the complete graph on $\mathbb{N}_{\geq 0}$ with weights $w(\{s, t\}) = \varepsilon(d)^{\|s-t\|}$ is a $C(d)$ -uniform weighted dependency graph for the σ_i 's.

→ deviation and speed of convergence estimates for the $\sum \sigma_i$.

(Non-uniform) weighted dependency graphs

(Non-uniform) weighted dependency graphs

Definition (Reminder)

Let C be a positive constant.

We say that G is a C -uniform weighted dependency graph for $\{Y_i, 1 \leq i \leq N\}$ if, for any i_1, \dots, i_r in $[N]$, we have

$$|\kappa(Y_{i_1}, \dots, Y_{i_r})| \leq C^r \sum_{T \text{ Cayley tree}} \prod_{\{j,k\} \in E_T} w(\{i_j, i_k\}), \quad (\text{UWDG})$$

(Non-uniform) weighted dependency graphs

Definition (F. 2016)

Let $\mathbf{C} = (C_r)_{r \geq 2}$ be a sequence of positive numbers.

We say that G is a \mathbf{C} -weighted dependency graph for $\{Y_i, 1 \leq i \leq N\}$ if, for any i_1, \dots, i_r in $[N]$, we have

$$|\kappa(Y_{i_1}, \dots, Y_{i_r})| \leq C_r \sum_{T \text{ Cayley tree}} \prod_{\{j,k\} \in E_T} w(\{i_j, i_k\}), \quad (\text{UWDG})$$

The bounds are now non-uniform in r .

(Non-uniform) weighted dependency graphs

Definition (F. 2016)

Let $\mathbf{C} = (C_r)_{r \geq 2}$ be a sequence of positive numbers.

We say that G is a \mathbf{C} -weighted dependency graph for $\{Y_i, 1 \leq i \leq N\}$ if, for any i_1, \dots, i_r in $[N]$, we have

$$|\kappa(Y_{i_1}, \dots, Y_{i_r})| \leq C_r \max_T \prod_{\{j,k\} \in E_T} w(\{i_j, i_k\}), \quad (\text{UWDG})$$

We can replace the sum by a max. Spanning trees of maximal weight can be found efficiently (e.g. Prim's algorithm).

(Non-uniform) weighted dependency graphs

Definition (F. 2016)

Let $\mathbf{C} = (C_r)_{r \geq 2}$ be a sequence of positive numbers.

We say that G is a \mathbf{C} -weighted dependency graph for $\{Y_i, 1 \leq i \leq N\}$ if, for any i_1, \dots, i_r in $[N]$, we have

$$|\kappa(Y_{i_1}, \dots, Y_{i_r})| \leq C_r \max_T \prod_{\{j,k\} \in E_T} w(\{i_j, i_k\}), \quad (\text{UWDG})$$

Lemma

$$|\kappa_r(\sum Y_i)| \leq C_r N D^{r-1}$$

Good to prove CLT, but not for mod-Gaussian convergence.

Applications of non-uniform weighted dependency graphs

- crossings in random pair-partitions;
- subgraph counts in $G(n, M)$;
- random permutations;
- particles in symmetric simple exclusion process;
- linear functional of Markov chain;
- spins in Ising model (with Jehanne Dousse);
- (*) patterns in multiset permutations, in set-partitions (with Marko Thiel);

*in progress. In blue: the ones which are also uniform WDG.

(Some of these applications use a variant of the above definition/lemma, which is more technical to state. . .)

Stability by powers of weighted dependency graphs

Setting:

- Let $\{Y_\alpha, \alpha \in A\}$ be r.v. with weighted dependency graph \tilde{L} ;
- fix an integer $m \geq 2$;
- for a multiset $B = \{\alpha_1, \dots, \alpha_m\}$ of elements of A , denote

$$\mathbf{Y}_B := Y_{\alpha_1} \cdots Y_{\alpha_m}.$$

Stability by powers of weighted dependency graphs

Setting:

- Let $\{Y_\alpha, \alpha \in A\}$ be r.v. with weighted dependency graph \tilde{L} ;
- fix an integer $m \geq 2$;
- for a multiset $B = \{\alpha_1, \dots, \alpha_m\}$ of elements of A , denote

$$\mathbf{Y}_B := Y_{\alpha_1} \cdots Y_{\alpha_m}.$$

Proposition (F., 2016)

The set of r.v. $\{\mathbf{Y}_B\}$ has a weighted dependency graph \tilde{L}^m , where

$$\text{wt}_{\tilde{L}^m}(\mathbf{Y}_B, \mathbf{Y}_{B'}) = \max_{\alpha \in B, \alpha' \in B'} \text{wt}_{\tilde{L}}(Y_\alpha, Y_{\alpha'}).$$

In short: if we have a dependency graph for some variables Y_α , we have also one for **monomials in the Y_α** .

☹ No analogue for uniform weighted dependency graphs.

Multilinear statistics in Markov chains

Setting:

- Let $(M_i)_{i \geq 0}$ be an irreducible aperiodic Markov chain on a finite space state S ;
- Assume M_0 is distributed with the stationary distribution π ;
- Set $Y_{i,s} = \mathbf{1}_{M_i=s}$.

Multilinear statistics in Markov chains

Setting:

- Let $(M_i)_{i \geq 0}$ be an irreducible aperiodic Markov chain on a finite space state S ;
- Assume M_0 is distributed with the stationary distribution π ;
- Set $Y_{i,s} = \mathbf{1}_{M_i=s}$.

Proposition

We have a weighted dependency graph \tilde{L} with $\text{wt}_{\tilde{L}}(\{Y_{i,s}, Y_{j,t}\}) = |\varepsilon|^{j-i}$.

We already know that: in fact this is a uniform weighted dependency graph.

Multilinear statistics in Markov chains

Setting:

- Let $(M_i)_{i \geq 0}$ be an irreducible aperiodic Markov chain on a finite space state S ;
- Assume M_0 is distributed with the stationary distribution π ;
- Set $Y_{i,s} = \mathbf{1}_{M_i=s}$.

Proposition

We have a weighted dependency graph \tilde{L} with $\text{wt}_{\tilde{L}}(\{Y_{i,s}, Y_{j,t}\}) = |\varepsilon|^{j-i}$.

Corollary (using the stability by product)

We have a weighted dependency graph \tilde{L}^m for monomials $Y_{i_1, s_1} \cdots Y_{i_m, s_m}$.

Multilinear statistics in Markov chains

Setting:

- Let $(M_i)_{i \geq 0}$ be an irreducible aperiodic Markov chain on a finite space state S ;
- Assume M_0 is distributed with the stationary distribution π ;
- Set $Y_{i,s} = \mathbf{1}_{M_i=s}$.

Proposition

We have a weighted dependency graph \tilde{L} with $\text{wt}_{\tilde{L}}(\{Y_{i,s}, Y_{j,t}\}) = |\varepsilon|^{j-i}$.

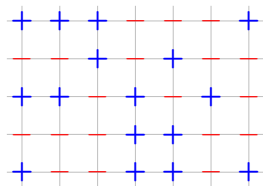
Corollary (using the stability by product)

We have a weighted dependency graph \tilde{L}^m for monomials $Y_{i_1, s_1} \cdots Y_{i_m, s_m}$.

→ gives a CLT for the number of copies of a given word in $(M_i)_{0 \leq i \leq N}$.
(Answers a question of Bourdon and Vallée.)

(But no uniform control on cumulants here 😞)

Back to Ising model



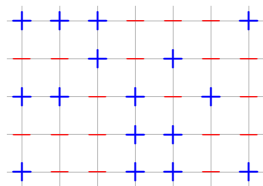
$$\mathbb{P}(\omega) \propto \exp [-H(\omega)];$$

$$H(\omega) = -\beta \sum_{x \sim y} \omega_x \omega_y - h \sum_x \omega_x.$$

Theorem

*In presence of a magnetic field or **at very low** or very large temperature, there exists $\varepsilon = \varepsilon(d, h, \beta) > 0$ such that the complete graph on \mathbb{Z}^d with weight $\varepsilon^{\|x-y\|_1}$ on the edge $\{x, y\}$ is a weighted dependency graph for $\{\sigma_x, x \in \mathbb{Z}^d\}$*

Back to Ising model



$$\mathbb{P}(\omega) \propto \exp [-H(\omega)];$$

$$H(\omega) = -\beta \sum_{x \sim y} \omega_x \omega_y - h \sum_x \omega_x.$$

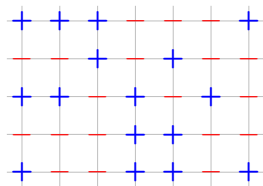
Theorem

*In presence of a magnetic field or **at very low** or very large temperature, there exists $\varepsilon = \varepsilon(d, h, \beta) > 0$ such that the complete graph on \mathbb{Z}^d with weight $\varepsilon^{\|x-y\|_1}$ on the edge $\{x, y\}$ is a weighted dependency graph for $\{\sigma_x, x \in \mathbb{Z}^d\}$*

Low temperature case proved by Malyshev and Minlos ('91).

Agian, proof based on cluster expansion. . .

Back to Ising model



$$\mathbb{P}(\omega) \propto \exp [-H(\omega)];$$

$$H(\omega) = -\beta \sum_{x \sim y} \omega_x \omega_y - h \sum_x \omega_x.$$

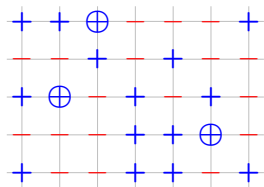
Theorem

*In presence of a magnetic field or **at very low** or very large temperature, there exists $\varepsilon = \varepsilon(d, h, \beta) > 0$ such that the complete graph on \mathbb{Z}^d with weight $\varepsilon^{\|x-y\|_1}$ on the edge $\{x, y\}$ is a weighted dependency graph for $\{\sigma_x, x \in \mathbb{Z}^d\}$*

Question: does it hold near the critical point?

(At the critical point, the answer is NO, since already covariances do not decay exponentially)

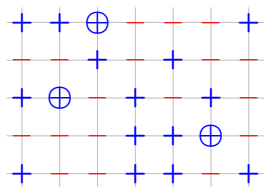
Ising model: CLT for global patterns



Circled spins:
occurrence of the + pattern 231

(notion inspired from patterns in permutations.)

Ising model: CLT for global patterns



Circled spins:
occurrence of the + pattern 2 3 1

$S_n^{\mathcal{P}}$:= number of occurrences of \mathcal{P} within $\Lambda_n = [-n, n]^d$.

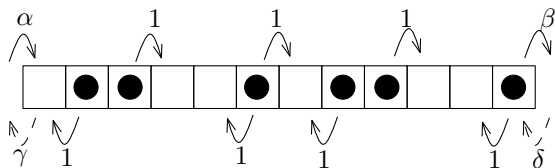
Theorem (Dousse, F., 2016)

Assume $\text{Var}(S_n^{\mathcal{P}}) \geq c \text{st} |\Lambda_n|^{2|\mathcal{P}|-2+\eta}$ for $\eta > 0$. Then we have

$$\frac{S_n^{\mathcal{P}} - \mathbb{E}(S_n^{\mathcal{P}})}{\sqrt{\text{Var}(S_n^{\mathcal{P}})}} \xrightarrow[n \rightarrow \infty]{d} \mathcal{N}(0, 1).$$

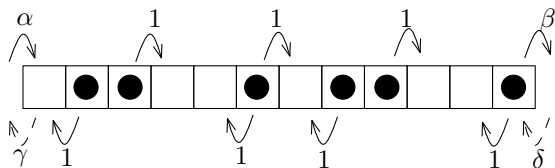
The lower bound of the variance is always fulfilled for patterns of only positive spins (as in the example).

Symmetric simple exclusion process (SSEP)



$\tau = (\tau_1, \dots, \tau_N)$ particle configuration with stationary distribution.

Symmetric simple exclusion process (SSEP)



$\tau = (\tau_1, \dots, \tau_N)$ particle configuration with stationary distribution.

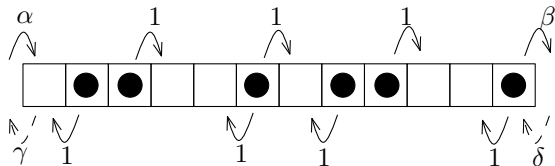
Theorem

The complete graph on $[N]$ with weight $1/N$ on each edge is a weighted dependency graph for the family $\{\tau_i, 1 \leq i \leq N\}$.

In particular, for disjoint i_1, \dots, i_r ,

$$\kappa(\tau_{i_1}, \dots, \tau_{i_r}) = \mathcal{O}_r(N^{-r+1}).$$

Symmetric simple exclusion process (SSEP)



$\tau = (\tau_1, \dots, \tau_N)$ particle configuration with stationary distribution.

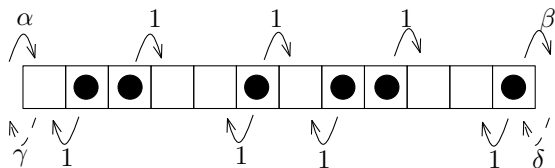
Theorem

The complete graph on $[N]$ with weight $1/N$ on each edge is a weighted dependency graph for the family $\{\tau_i, 1 \leq i \leq N\}$.

Ingredients of the proof:

- joint moments of the τ_i given by matrix ansatz;
- in case of SSEP, this gives an induction formula for cumulants (Derrida, Lebowitz, Speer, 2006).

Symmetric simple exclusion process (SSEP)



$\tau = (\tau_1, \dots, \tau_N)$ particle configuration with stationary distribution.

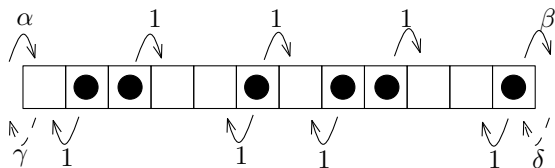
Theorem

The complete graph on $[N]$ with weight $1/N$ on each edge is a weighted dependency graph for the family $\{\tau_i, 1 \leq i \leq N\}$.

Consequences:

- functional CLT for the particle distribution function $F_\tau(u) = \sum_{i=1}^{\lfloor uN \rfloor} \tau_i$;
- also, e.g., for the number $\sum_i \tau_i(1 - \tau_{i+1})$ of particles that can jump to their right (using stability by powers).

Symmetric simple exclusion process (SSEP)



$\tau = (\tau_1, \dots, \tau_N)$ particle configuration with stationary distribution.

Theorem

The complete graph on $[N]$ with weight $1/N$ on each edge is a weighted dependency graph for the family $\{\tau_i, 1 \leq i \leq N\}$.

The same is conjectured for ASEP in general.

Conclusion

- **Dependency graphs provide uniform bounds on cumulants** (lots of examples, including subgraph counts in $G(n, p)$ for fixed p , pattern occurrences in random permutations);

Conclusion

- **Dependency graphs provide uniform bounds on cumulants** (lots of examples, including subgraph counts in $G(n, p)$ for fixed p , pattern occurrences in random permutations);
- There is a **weighted variant** which include linear statistics of Markov chain, magnetization in Ising model (except at low temperature);

Conclusion

- **Dependency graphs provide uniform bounds on cumulants** (lots of examples, including subgraph counts in $G(n, p)$ for fixed p , pattern occurrences in random permutations);
- There is a **weighted variant** which include linear statistics of Markov chain, magnetization in Ising model (except at low temperature);
- If you **relax assumptions** to only **get a CLT** and not a uniform control on cumulants, the weighted version applies to **even more examples**;

Conclusion

- **Dependency graphs provide uniform bounds on cumulants** (lots of examples, including subgraph counts in $G(n, p)$ for fixed p , pattern occurrences in random permutations);
- There is a **weighted variant** which include linear statistics of Markov chain, magnetization in Ising model (except at low temperature);
- If you **relax assumptions** to only **get a CLT** and not a uniform control on cumulants, the weighted version applies to **even more examples**;
- We would really like to get uniform bounds on cumulants for these extra examples. . .